

HAWKES BINOMIAL TOPIC MODEL WITH APPLICATIONS TO COUPLED CONFLICT-TWITTER DATA

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Abstract We consider the problem of modeling and clustering heterogeneous event data arising from coupled conflict event and social media data sets. In this setting conflict events trigger responses on social media and at the same time signals of grievance detected in social media may serve as leading indicators for subsequent conflict events. For this purpose we introduce the Hawkes Binomial Topic Model (HBTM) where marks, Tweets, and conflict event descriptions are represented as bags of words following a Binomial distribution. When viewed as a branching process, the daughter event bag of words is generated by randomly turning on/off parent words through independent Bernoulli random variables. We then use Expectation-Maximization to estimate the model parameters and branching structure of the process. The inferred branching structure is then used for Topic cascade detection, short-term forecasting, and investigating the causal dependence of grievance on social media and conflict events in recent elections in Nigeria and Kenya.

1. Background and Motivation. Twitter and other social media platforms have emerged as important tools for the public to communicate responses to crises and terrorist attacks [5] and more generally to communicate collectively, exchanging grievances that may catalyze mobilization [3, 29]. Research has focused on understanding public sentiment around these types of events and determining the central actors in the social network that are key to shaping public response [5] along with measuring short-term changes in the intensity of conflict using social media [37] or considering effects from regional instability [4].

At a more macro spatial-temporal scale, recent research has focused on modeling the endogenous and exogenous processes that generate thousands of terrorist and conflict events at the level of countries and years or decades. For this purpose point processes are used to model contagion effects in the risk of terrorist activity [26] and both contagion and exogenous rate fluctuations in conflict [21, 36, 35]. Because of the relative infrequency of conflict and terrorist events, having auxiliary data that can provide a signal for the risk of future events is highly desirable. While conflict events have been shown to influence overall regional instability [4], point process models of conflict to date have focused on univariate data [26, 21, 36, 35]. In this paper we use

Twitter posts relevant to elections as auxiliary data to investigate whether we can provide a leading indicator for conflict activity and contentious events. Because only a subset of Tweets may be relevant to a particular event, we introduce a Hawkes topic model to capture Tweet content and conflict event descriptions.

We propose a Hawkes Binomial Topic Model (HBTM) where marks (Tweets and conflict event descriptions) are represented as bags of words following a Binomial distribution. When viewed as a branching process, the daughter event bag of words is generated by randomly turning on/off parent words through independent Bernoulli random variables. We allow for a secondary

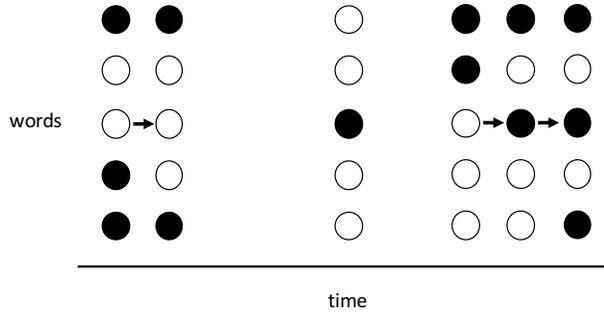


FIGURE 1. In the HBTM, spontaneous events occur with marks generated by a binomial random variable over the dictionary of keywords contained in the data set. Events then trigger offspring events whose marks are generated by switching parent event words off with probability p_{off} and on with probability p_{on} .

mark, representing event type (Twitter or conflict event type) and extend the HBTM to a mutually exciting point process that captures cross-excitation effects between different types of conflict data, along with social media data relevant to the intensity of conflict. We then analyze a unique merged data set of all conflict events from the Armed Conflict and Location Event Data Project, or ACLED [27] from a 90 day period around national-level election votes, along with a data set of tweets extracted using a keyword relevancy algorithm about the election during the same time period. We perform this analysis for the 2015 Nigeria and 2013 Kenya presidential elections.

We select the Kenya and Nigeria elections because they are highly likely to contain ethnic and religious grievances and contain varying levels of contentious activity, described in detail in Section 3.1. Grievances, or collective emotional responses to perceived injustices, are theorized to decrease barriers to mobilization by focusing frustrations into aggression [14]. We expect

the emotions generated by underlying grievances to be expressed in election-relevant communication on social media with more volatility over time than the inequalities that generate them. We capture the emotive content as contentious topics discussed before, during, and after the vote to activate salient group divisions underlying political cleavages expecting to win or lose, fairly or unfairly, with the incumbent or opposition. We test whether variation in this communication, through topics and keywords, may provide a leading indicator for violent and non-violent events. We find the mutually exciting HBTM accounts for topic spatial and temporal contagion of risk between conflict events of different types and Tweets related to the elections on social media.

The output of the model is an estimated intensity of events along with identified event cascades across topic space and time. Our approach is related to stochastic declustering [40] of earthquake catalogs, however our objective is to detect clusters and measure mutual excitation, whereas in seismology the goal is to remove clusters to obtain an estimate of the intensity of main shock earthquakes. Our general methodology proceeds by viewing the point process as a branching process and then uses Expectation-Maximization (EM) to estimate the causal dependency of events and detect cascades of related events.

Previous studies on Hawkes models of social media have utilized temporal point processes [38, 30], Latent Dirichlet Allocation (LDA) [19] and Correlated Topic Model (CTM)-based point processes [15]. Dirichlet Hawkes process is introduced in [8, 34] for modeling social media and large-scale real-time burst modeling on Twitter is considered in [33]. In [11], a joint model for information diffusion and an evolving network is developed. In [9], the authors use Hawkes topic modeling for detecting fake retweeters and in [39] the authors use Hawkes topic modeling to analyze the dark web. Latent influencers are modeled in [31] using an Indian buffet Hawkes process. For a review of point process modeling of social diffusion see [16].

Our work here considers a novel application, namely detecting mutual excitation between online Twitter activity and real-world conflict and political instability events. The model we introduce is unique both in terms of the mutually-exciting component of the model capturing effects across heterogeneous data sets and the Binomial Topic Model that facilitates a fast, easy-to-implement EM algorithm for inference while still outperforming the LDA Hawkes process used in [19]. Overall, the contributions of this work are as follows:

- We analyze a novel data set of election related tweets along with event data on protests, violence, and territory change and estimate the casual

influence across social media and conflict.

- We introduce a novel mutually exciting Hawkes Binomial Topic Model. The mutual excitation component is novel for merging heterogeneous event-text data sets and estimating causality across them. The HBTM is novel as a model for co-occurrence of words between parent-child events in the Hawkes process. We show improved coherence of HBTM over LDA.

The outline of the paper is as follows. In Section 2 we introduce the Hawkes Topic Binomial Model and its extension to mutual excitation. Section 3 describes the Nigeria and Kenya elections, the conflict event data, and the methods used to retrieve and identify relevant Tweets. In Section 4 we present inference results on synthetic data and then study the detection of event cascades across heterogeneous event types and the overall causal structure of the estimated model applied to the social media and ACLED datasets. We assess the model’s ability to forecast the level of conflict one day ahead against point process models with only univariate data as input and models without Topic marks. In Section 5 we discuss the significance of our work along with future research directions.

2. Hawkes Binomial Topic Model. Consider a marked Hawkes process with intensity $\lambda(t, \vec{m})$ determined by,

$$(1) \quad \lambda(t, \vec{m}) = \mu J_0(\vec{m}|p_0) + \sum_{t > t_i} \theta \omega e^{-\omega(t-t_i)} J_1(\vec{m}, \vec{m}_i | p_{off}, p_{on}).$$

Here events at time t_i are associated with a mark \vec{m}_i , a vector of size W , the number of words in the overall dictionary across events. The binary variables indicate whether each word is present or absent in the event at time t_i . Spontaneous events occur according to a Poisson process with rate μ . The mark vector of spontaneous events is determined by,

$$(2) \quad J_0(\vec{m}|p_0) = p_0^{\sum_{j=1}^W m_j} (1 - p_0)^{W - \sum_{j=1}^W m_j},$$

which is the product of W independent Bernoulli random variables with parameters p_0 .

The parameter θ determines the expected number of events triggered by each event and the expected waiting time between a parent-daughter event pair is given by ω^{-1} . The mark of a daughter event is determined by two independent Bernoulli processes. Each word absent, or “turned off,” in the parent bag of words is added to the bag of words of the child event with

probability p_{on} . Each word present in the parent bag of words is deleted with probability p_{off} . Thus J_1 is given by,

$$(3) \quad J_1(\vec{m}, \vec{m}_i | p_{off}, p_{on}) = p_{on}^{W_1^{\vec{m}, \vec{m}_i}} (1 - p_{on})^{W_2^{\vec{m}, \vec{m}_i}} p_{off}^{W_3^{\vec{m}, \vec{m}_i}} (1 - p_{off})^{W_4^{\vec{m}, \vec{m}_i}},$$

where $W_1^{\vec{m}, \vec{m}_i}$ is the number of words present in the child vector and absent in the parent vector, $W_2^{\vec{m}, \vec{m}_i}$ is the number of words absent in both vectors, $W_3^{\vec{m}, \vec{m}_i}$ is the number of words in the parent vector absent in the child vector, and $W_4^{\vec{m}, \vec{m}_i}$ is the number of words present in both vectors.

While this model specification may not be appropriate for all dynamic topic analyses, we believe it works well when documents are short sentences or bags of words, as is the case with Tweets and event descriptions. After removing stop words we restrict the dictionary to the W most frequent words, on the order of several hundred most frequent words across events.

2.1. *Estimation Via Expectation-Maximization.* The Model given by Eq. 1 can be viewed as a branching process [22][32] where events occur according to a stationary Poisson process $\mu J_0(\vec{m} | p_0)$ and then each event i generates a Poisson process with intensity $\theta \omega e^{-\omega(t-t_i)} J_1(\vec{m}, \vec{m}_i | p_{off}, p_{on})$. Let $u_{ij} = 1$ when event i is the direct offspring of event j and 0 otherwise and $u_i^b = 1$ when event i is a ‘‘background’’ or ‘‘spontaneous’’ event generated by the background Poisson process (and 0 otherwise). Given knowledge of u_{ij} , the estimation problem decouples into several independent Poisson estimation problems. However, because u_{ij} is unknown we introduce a matrix p_{ij} representing the probability that event j triggered event i and a vector p_i^b representing the probability that event i is a background event. Expectation-maximization inference then proceeds by iterating between the E-step:

$$(4) \quad p_{ij} = \frac{\theta \omega e^{-\omega(t_i-t_j)} J_1(\vec{m}_i, \vec{m}_j | p_{off}, p_{on})}{\lambda(t_i, \vec{m}_i)}$$

$$(5) \quad p_i^b = \frac{\mu J_0(\vec{m}_i | p_0)}{\lambda(t_i, \vec{m}_i)}$$

and the M-step:

$$(6) \quad \omega = \frac{\sum_{t_i > t_j} p_{ij}}{\sum_{t_i > t_j} p_{ij}(t_i - t_j)}$$

$$(7) \quad \mu = \frac{\sum_{i=1}^N p_i^b}{T}$$

$$(8) \quad \theta = \frac{\sum_{t_i > t_j} p_{ij}}{N}$$

$$(9) \quad p_0 = \frac{\sum_{i=1}^N p_i^b \|\vec{m}_i\|_1 / W}{\sum_{i=1}^N p_i^b}$$

$$(10) \quad p_{on} = \frac{\sum_{t_i > t_j} p_{ij} W_1^{\vec{m}_i, \vec{m}_j} / (W_1^{\vec{m}_i, \vec{m}_j} + W_2^{\vec{m}_i, \vec{m}_j})}{\sum_{t_i > t_j} p_{ij}}$$

$$(11) \quad p_{off} = \frac{\sum_{t_i > t_j} p_{ij} W_3^{\vec{m}_i, \vec{m}_j} / (W_3^{\vec{m}_i, \vec{m}_j} + W_4^{\vec{m}_i, \vec{m}_j})}{\sum_{t_i > t_j} p_{ij}}$$

Here N is the number of events in the data set and T is the length of the time window. We note that the choice of the triggering kernel allows the parameters to be determined in the M-step via weighted sample mean estimators [20]. In Equations 4-11 we have made the approximation that the integrals of the triggering kernel defined over the observation window can be replaced by integrals over the entire space [20].

2.2. Mutually Exciting HBTM. Next we introduce an extension of HBTM allowing for mutual excitation across heterogenous data sets. For example, in Section 4 we model the coupled system of ACLED events, with the subcategories protests, violence against civilians, and territory change, along with tweets from the 90 day period around an election yielding four total event categories. Twitter and conflict activities operate on different timescales, so we wish to allow for different Hawkes parameter values for both self and mutual excitation.

For this purpose we add a second mark variable, $s = 1, \dots, S$, indicating the category of each event out of S categories. The conditional intensity is then determined by,

$$(12) \quad \lambda_s(t, \vec{m}) = \mu_s J_0(\vec{m} | p_0^s) + \sum_{t_i > t_j} \theta_{ss_i} \omega_{ss_i} e^{-\omega_{ss_i}(t-t_i)} J_1(\vec{m}, \vec{m}_i | p_{off}^{ss_i}, p_{on}^{ss_i}).$$

Here events of type s occur either according to a temporally stationary Poisson process with intensity $\mu_s J_0(\vec{m} | p_0^s)$ or may be instead triggered by a previous event in the history of the process. The model parameters μ_s , p_0^s , θ_{sv} , ω_{sv} , p_{off}^{sv} , and p_{on}^{sv} now depend on $s = 1, \dots, S$ and $v = 1, \dots, S$ the categories of the child and parent events respectively. The EM algorithm for Equation 12 is analogous to Equation 4-11. Here the only difference is that the summations over $t_i > t_j$ for parameter with index s, v are restricted to i, j such that $s_i = s, s_j = v$.

3. Cases, Data, and Methods for Retrieval. In this section we describe our case selection of the 2013 Kenya and 2015 Nigeria elections and the conflict event data from ACLED. We then summarize our data collection process for retrieving Tweets that contain politically relevant communication. Expressions of grievance may be captured in this political content, thereby coupling social media messages to violent or contentious events.

3.1. *Case Selection.* To investigate causation between social media and violent political events, we must select cases that exhibit two properties: high levels of social media use, and a variety of events with varying degrees of violence. Sub-Saharan Africa presents several cases that satisfy these constraints and has wide variation in infrastructure, social media use, and political violence, so we focus our research on this geographical area. To capture political discourse, we further concentrate our work on timeframes surrounding election events in Sub-Saharan Africa, as these events generally promote political discussion and increase the overall volume of discourse online. For this specific research, we study the 2015 Nigerian presidential election and the 2013 Kenyan presidential election, as both countries have substantial Internet-using populations and numerous relevant events. Regarding Internet use, 45.5% of Nigerians and 33.5% of Kenyans report using the Internet at least a few times a month, a few times a week, or daily[2]. Based on qualitative examination, we expect the existence of underlying religious grievances in Nigeria and ethnic grievances in Kenya to impact variation in the number of violent and non-violent events over time and space. Nigeria’s 2015 election contains 565 events, and Kenya’s 2013 election contains 281.

Below, we provide a brief summary of these elections to provide context for this study.

3.1.1. *2015 Nigerian Presidential Election.* On March 28, 2015, 42% of Nigeria turned out to vote in Nigeria’s presidential election, pitting incumbent Goodluck Jonathan against Muhammadu Buhari, a retired general from the North with a strong anti-corruption platform. This election was a near-identical re-enactment of the 2011 electoral showdown between these two candidates, but this time, Buhari succeeded in unseating Jonathan. This opposition win made history in Nigeria as the first time an incumbent was defeated and peacefully stepped down [24]. Despite the peaceful transition, the election was emotionally charged on both sides as the country battled security issues and persistent accusations of electoral manipulation. State forces were too occupied with Boko Haram to ensure safe voting, resulting in low voter turnout, likely owing to anticipated violence. The incumbent administration also postponed the election, citing the security situation, serving only

to inflame accusations of chicanery. The opposition accused the incumbent of manipulating the timing to advantage his party over challenger. The mixture of inaction and dysfunction on the security question likely led many to believe that Buhari was a better choice than incumbent Jonathan. While in Nigeria, voter preferences tend to fall along ethnic and religious lines, parties cannot win without forming alliances and gaining support from outside their dominant religious or ethnic groups. As such, enough southern Christians voted for Buhari to secure a solid victory.

3.1.2. *2013 Kenyan Presidential Election.* On March 4, 2013, 86% of Kenyans turned out to vote in the presidential election, the first since catastrophic levels of violence occurred in 2007. In the previous election in 2007, at least 1,133 people perished, and 700,000 people were displaced [7]. In 2013, however, Kenya’s ethnic grievances responsible for violence in 2007 were dampened through considerable efforts at preventing violence, including a joint ticket between ethnic rivals Uhuru Kenyatta and William Ruto. Kenyatta, the incumbent from 2007’s election, and former-rival-turned-running-mate William Ruto united to form the new umbrella Jubilee Coalition. While the International Criminal Court had individually indicted both Kenyatta and Ruto for their roles in the 2007 electoral violence, their alliance mitigated potential violence by aligning the rivaling Kalenjin and Kikuyu communities at the heart of the 2007 violence. Running against Kenyatta and Ruto was Raila Odinga and running mate Kalonzo Musyoka under the umbrella Coalition for Reforms and Democracy. All candidates regularly reiterated their commitment to peace, media outlets were trained in “conflict-sensitive reporting” in order to avoid inflaming tensions, and security forces were heavily deployed in sensitive areas [7]. This new election was considered free, fair and peaceful.

3.2. *Overview of Conflict Event Data.* ACLED contains the broadest range of diverse event types of data sets that track conflict events [10], focusing on civil and communal conflicts, violence against civilians, remote violence, and riots and protests [28]. The ACLED project codes the dates and locations of all reported protests and political violence within periods of instability, civil war, and regime breakdown, capturing episodic activity, such as potential precursor events. Actors include political agents, including governments, rebels, militias, communal groups, political parties, rioters, protesters and civilians. From these event types, we utilize protests, violence against civilians, and transfer of territory to cover the broad spectrum of activity. Using the ACLED’s date, location, event type, and event description text we create a unique data set by merging the observations with Tweets

for the 90 day period around the election. These Tweets are curated using the following techniques to retrieve information, for query expansion and for identification of relevance within the corpus and metadata.

3.3. *Social Media Data Collection.* To study the connection between social media and conflict, we must collect large volumes of relevant social media. The Twitter platform is ideal for this task as its content is open and publicly available by default (in contrast to Facebook or Snapchat, which restrict access to pre-selected individuals by default). Nigeria and Kenya also are in the top 25% of countries with presence on Twitter in terms of Tweets with geospatial location tags, called "geotags" and communicate primarily in English on Twitter. This section details our process for retrieving relevant content from Twitter's platform, mitigating biases in this content, and adapting to "big data"-scale datasets. At a high level, this retrieval process consists of:

1. Human-expert query generation,
2. Computationally assisted, multi-dimensional query expansion,
3. Data reseller-backed retrieval, and
4. At-scale data sanitization.

3.3.1. *Expert Query Generation.* To prime our data collection pipeline, we ask subject-matter experts who have studied the Nigerian and Kenyan elections to generate semi-structured descriptions of each election. These descriptions include collections of keywords, phrases, and individuals for which we can search in a social media dataset. Search terms could be lists of individuals involved, such as leaders, candidates, journalists, and activists; associated organizations, supporting constituencies, and ideologies; issues or drivers leading to contentious actions; and actions or events in the relevant time period, such as protests, rallies, or attacks, court cases, voter registration problems, candidate or party announcements, scandals, or political violence. Experts then distill these summaries into topics and sub-topics and refine their descriptions through an iterative search process that uses both social and traditional media sources.

3.3.2. *Multi-Dimensional Query Expansion.* Though experts can generate high-quality event descriptions, the Internet's varied populations and communities may express thoughts or refer to the same concepts in many different ways. We therefore assume event descriptions are incomplete and may omit a subset of tokens some populations may use to refer to a particular entity or concept. To address this issue, we leverage a large, undirected

sample of Twitter data to expand these expert-generated queries and capture relevant tokens in a data-driven manner. This sample is collected from Twitter’s public sample stream, which publishes 1% of all Tweets posted to the Twitter platform. Furthermore, “subtweeting,” or referencing an event or entity without directly mentioning it, complicates retrieval, so we also expand queries using social and spatial dimensions (the spatial dimension is expanded via the data reseller).

Textual expansion is performed by querying our undirected Twitter sample for content that matches expert-generated queries. We then identify keywords whose frequencies in the set of matching messages are much higher than in the general sample by ranking tokens by their Kullback-Leibler (KL) divergence [17]. Keywords with a KL divergence exhibit a strong co-occurrence connection with the original query but are rare in the underlying sample. These keywords are useful for identifying potentially missing keywords for the event particular to Twitter and its communities’ differing vocabularies [13]. Specifically, we find all messages in the random sample that match the original query, tokenize these messages into bags of words, rank by KL divergence score, and ask our experts to determine which keywords to add to the original query.

To expand queries along the social dimension, we identify socially-relevant users by converting the retweet and mention activity in a sample of relevant content into a directed graph of interactions. This relevant sample is generated from Tweets matching the original expert queries described above. Research shows highly followed or retweeted users are often not the most influential users [6], so we follow Kwak et al. and use a version of Google’s PageRank algorithm to calculate an “authority” score for users in this interaction graph [18]. Experts are then presented with a list of accounts ranked by their authority score and asked which accounts should be added to the original queries.

3.3.3. Information Retrieval from Data Resellers. After expanding experts’ event descriptions and queries, we search Twitter’s full historical archive to acquire a large dataset of Tweets for each election. Twitter’s historical archive is not freely available, however, so we contract with Gnip, Twitter’s primary data reseller to access this data. We use Gnip’s native support for textual, social, spatial, and temporal queries to search for relevant Tweets that match keywords/phrases, mention/are from relevant accounts, or are posted from within the target country, all of which are restricted to 60 days prior and 30 days following the election date in each country.

Using this process, we retrieved 7.1 and 4.7 million Tweets for Nigeria and

Kenya respectively.

3.3.4. *At-Scale Data Sanitization.* The previous steps in this data retrieval process ensure significant *recall* in our data; we therefore avoid omitting relevant content (high recall) at the cost of including irrelevant Tweets (low precision). To ensure high-quality modeling with HBTM, however, we must clean these datasets and remove content that is not relevant to the political contexts of these countries. Standard approaches have experts sanitize the data through review and ensuring relevance, but when dealing with several million social media messages, complete human review is intractable. To support sanitization at scale, we use a mixed-methods, iterative approach to train an automated system that can identify relevant Tweets.

We bootstrap this sanitization process by randomly sampling 300 Tweets that match textual, social, and spatial queries, resulting in a set of approximately 900 Tweets. We then perform several rounds of relevance feedback in which pairs of expert analysts review these 300 tweets, labeling each as relevant, irrelevant, not English, or undecidable. Labels are analyzed for inter-rater reliability, experts meet to discuss their labeling strategies, and labeling is repeated until agreement between experts reaches a sufficient level of agreement, which we set as a Cohen’s $\kappa > 0.61$. Tweets agreed to be relevant and irrelevant are saved for classifier training, while undecided or non-English Tweets are discarded.

For Nigeria, our annotators achieved an agreement of $\kappa = 0.75$ and agreed on 836 labeled Tweets, 495 of which were labeled as relevant to the Nigerian election. In Kenya, agreement reached $\kappa = 0.69$ and 688 usable tweets, 321 of which were labeled as relevant.

To train a classifier to identify relevance from these datasets, we featurize each Tweet into a bag of words and train a set of Gradient Boosted Trees (GBTs)¹ on these feature vectors and their relevance labels. We train separate GBTs for each election event as well. We then apply these trained classifiers to the full set of Tweets for each event and store all instances classified as relevant. We then randomly sample 1,000 Tweets from these potentially relevant Tweets and pass them through a second round of relevance feedback to ensure our classifiers are performing well. Labels from this second round of relevance feedback are incorporated into our relevance models as well.

In this second round of assessment, our experts agreed 93% of the Nigerian

¹While other classifiers such as support vector machines and random forests perform equivalently (we tested these, and they achieved similar performance scores), GBTs are easily distributed across cluster systems.

Tweets the classifier identified as relevant were correct. For Kenya, however, our experts only found 49% of the 1,000 sampled Tweets to be relevant. These results suggest Nigeria has a large quantity of relevant data, but Kenya’s data is either noisy or our models cannot capture political discourse using only textual data.

To determine whether our focus on textual content was adequate for relevance classification, we ran an experiment that trained classifiers on the other dimensions present in Tweets as well: the standard bag-of-words model, a bag-of-mentions model (where tokens are replaced by Twitter users mentioned in a Tweet), bag-of-locations (using self-reported user locations), and temporal classification as well (i.e., how close is the Tweet to the event). For Nigeria and Kenya, we generated precision-recall curves for each set of features using 30 rounds of random sampling in which 25% of our labels were held out for testing, and a classifier was trained on the remaining 75%. Precision and recall were calculated across these 30 rounds, and we calculated the mean area under the curve (AUC) and 95% confidence intervals for each feature set and a union of all feature sets. Results, shown in Figure 2, show that textual features perform as well or better than all other features separately and together.

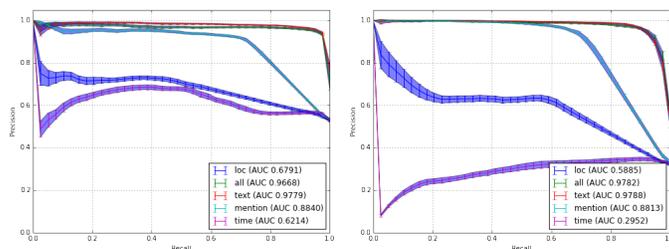


FIGURE 2. *Precision-Recall Curves for Various Dimensions in Nigeria (left) and Kenya (right)*

This sanitization process’s final step is to restrict data to only those Tweets classified as relevant *and* are from social media users who participate in communities discussing these election events. We identify communities in Tweets by applying label propagation to the interaction graph described above and keep the three largest communities who reference these events. This step is primarily useful to remove references to elections outside the target countries, as elections in the United States and India use similar language and hashtags that could contaminate our data.

This process yielded 2.8 million Tweets relevant to Nigeria and 1.1 million Tweets relevant to Kenya.

4. Results.

4.1. *Simulated Data.* We first test the HBTM EM estimation procedure by simulating several realizations of the HBTM process and assessing parameter recovery. In particular, we simulate a mutually exciting HBTM with two categories and parameters given by $\mu = [1, 2]$, $\theta = \begin{bmatrix} .5 & .1 \\ .2 & .7 \end{bmatrix}$, $\omega = \begin{bmatrix} .1 & 1 \\ .05 & .9 \end{bmatrix}$, $p_{on} = \begin{bmatrix} .05 & .1 \\ .05 & .07 \end{bmatrix}$, $p_{off} = \begin{bmatrix} .2 & .1 \\ .3 & .1 \end{bmatrix}$, $p_0 = [.1, .08]$, and $W = 200$.

In Figure 3 we plot the true value of the parameters along with a histogram of estimated parameters from 100 realizations of the point process where 10 iterations of EM were used in each simulation. Here we find that all parameters are recovered with error on the order of or less than 10%. We have posted the Matlab code to replicate the simulations on Github [1].

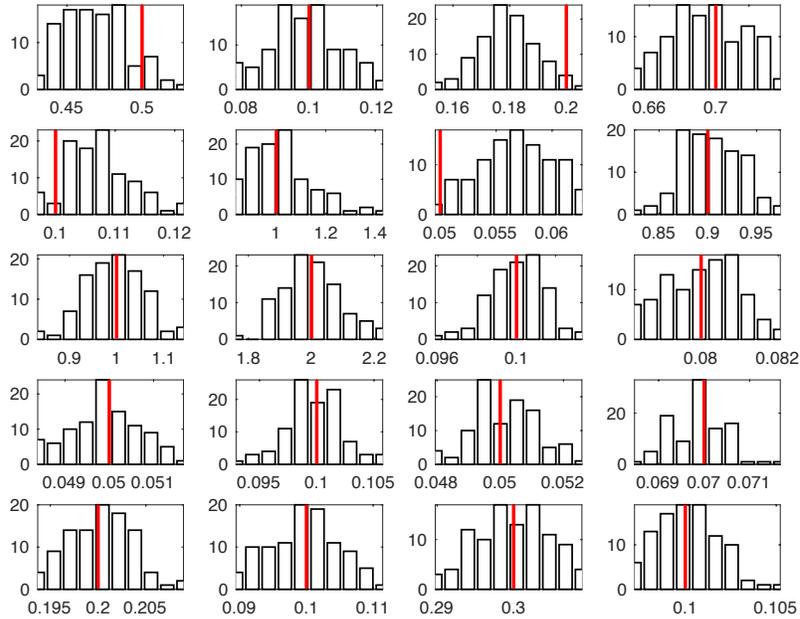


FIGURE 3. EM estimates for 100 realizations of the mutually exciting HBTM. Red line indicates ground truth. Top row corresponds to θ , second row corresponds to ω , third row corresponds to μ and p_0 , fourth row corresponds to p_{on} and fifth row corresponds to p_{off} .

4.2. *Coupled ACLED-Twitter Data.* We next analyze the merged ACLED and Twitter data set of events occurring in Nigeria from February 2 through April 30 2015 and Kenya from January 8 2013 to April 2 2013. In particular, we consider four categories of events in the given time periods: protests, violence against civilians, battles with territory change, and tweets relevant to the elections during the time periods. Given the quadratic cost of the EM algorithm we subsample the Twitter data after the relevancy algorithm is applied to bring the size of the Twitter data down to $O(10^4)$ events. In particular, the Kenya dataset has 6108 events and the Nigeria dataset has 8412 events.

We estimate the mutually exciting HBTM using the Expectation-Maximization algorithm from Section 2 and find that the EM algorithm converges within 10-20 iterations. We first investigate the extent of self and mutual excitation across events of different categories by constructing a weighted, directed graph where the edge weights are determined by $\theta_{ml}N_l$, representing the number of direct offspring events of category m generated by events of type category l . We include the background rates as nodes on the network as well, letting the edge weights be determined by $\mu_m T$, representing the number of background events of type m .

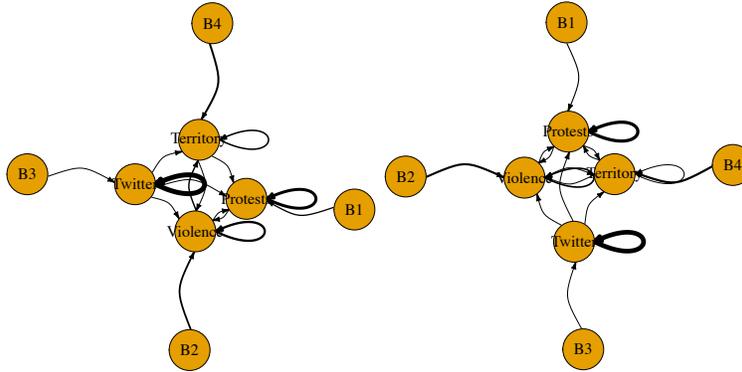


FIGURE 4. Graphical representation of estimated point process model fit to the ACLED-Twitter event data set for Nigeria (left) and Kenya (right). Weights of the edges of the directed graph correspond to the fraction of events triggered across the edge. B_1 - B_4 represent the background rate of the process.

Several patterns emerge upon inspection of the graph in Figure 4. First, in both Nigeria and Kenya Tweets are predominately generated through self-excitation. The lack of influence of ACLED events on Twitter may be in part explained by the lack of hourly resolution of ACLED events and the fast time

scale on which Tweet patterns evolve. Nigeria and Kenya graphs both have similar edge characteristics, with the exception of protest-territory mutual excitation. In Table 2-3 we display the parameter estimates for the ME-HBTM corresponding to the Graph in Figure 4. We obtain compute standard errors by repeated simulation of the process with the MLE parameters and computing bootstrapped estimates. In Table 3 we observe the productivity parameter values θ . For example, we estimate that in Kenya and Nigeria 1 protest is generated by every 2000 election related tweets ($\theta_{ij} = .0005$). Based on the standard errors this value is statistically significant.

TABLE 1
Mean and standard error for μ in ME-HBTM.

	Protest	Violence	Twitter	Territory
Nigeria	0.053 +/- 0.004	0.047 +/- 0.003	0.435 +/- 0.026	0.032 +/- 0.002
Kenya	0.056 +/- 0.008	0.080 +/- 0.011	4.108 +/- 0.108	0.001 +/- 0.003

We next investigate the extent to which the model predicts daily risk of ACLED event activity in the case of the mutually exciting HBTM (ME-HBTM), HBTM without mutual excitation (four independent Hawkes processes with θ diagonal), and a mutually exciting Hawkes process with no Topic mark component. For each day in the 90 day periods we average the intensity over the empirical mark distribution for that day and use the mean intensity to predict the number of events the following day. We then use the Normalized Gini Index (NGI) to assess each models performance [12]. The NGI is a continuous analog to AUC measuring the ability of each model to correctly sort days by the number of events, where errors for high event days are penalized greater than low event days (NGI is normalized such that 1

TABLE 2
Mean and standard error for ω in ME-HBTM.

Nigeria	Protest	Violence	Twitter	Territory
Protest	0.306 +/- 0.278	0.215 +/- 1.375	0.271 +/- 0.744	0.3989 +/- 4.620
Violence	0.776 +/- 2.185	0.255 +/- 0.054	1.360 +/- 0.318	0.250 +/- 0.152
Twitter	0.229 +/- 0.622	18.483 +/- 1.113	0.818 +/- 0.005	1.846 +/- 0.583
Territory	0.401 +/- 0.511	0.252 +/- 1.544	1.481 +/- 0.523	1.498 +/- 0.284
Kenya	Protest	Violence	Twitter	Territory
Protest	0.359 +/- 19.014	1.097 +/- 0.154	11.378 +/- 0.846	0.414 +/- 1.157
Violence	0.865 +/- 0.402	0.261 +/- 0.215	0.556 +/- 0.052	0.265 +/- 0.671
Twitter	21.706 +/- 2.025	0.699 +/- 0.220	2.497 +/- 0.007	3.612 +/- 2.933
Territory	1.360 +/- 3.716	0.197 +/- 6.042	0.541 +/- 0.659	0.217 +/- 0.125

TABLE 3
Mean and standard error for θ in ME-HBTM.

Nigeria	Protest	Violence	Twitter	Territory
Protest	0.697 +/- 0.032	0.062 +/- 0.017	0.025 +/- 0.016	0.002 +/- 0.006
Violence	0.079 +/- 0.012	0.561 +/- 0.023	0.008 +/- 0.019	0.322 +/- 0.020
Twitter	0.0005 +/- 0.0003	0.0004 +/- 0.0001	0.994 +/- 0.004	0.0001 +/- 0.0001
Territory	0.129 +/- 0.013	0.174 +/- 0.019	0.009 +/- 0.039	0.434 +/- 0.029
Kenya	Protest	Violence	Twitter	Territory
Protest	0.751 +/- 0.027	0.073 +/- 0.012	0.027 +/- 0.067	0.085 +/- 0.009
Violence	0.065 +/- 0.017	0.439 +/- 0.020	0.002 +/- 0.093	0.079 +/- 0.009
Twitter	0.0005 +/- 0.00006	0.0005 +/- 0.00005	0.940 +/- 0.002	0.0006 +/- 0.00006
Territory	0.184 +/- 0.033	0.591 +/- 0.043	0.104 +/- 0.307	0.299 +/- 0.020

TABLE 4
NGI scores for Kenya

	ME-HBTM	HBTM	ME-H
Protest	0.3990	0.3546	0.3558
Violence	0.2247	0.2357	0.1399
Territory	0.5798	0.6134	0.4958

indicates a perfect prediction).

In Table 4 and 5 we display the NGI scores for the three models. In both Kenya and Nigeria, protests are most accurately predicted using the ME-HBTM, likely due to protests being the most relevant ACLED event to the elections and social media discussions. In Kenya, the HBTM self-exciting model performs best for violence and territory event types. In Nigeria, territory is most accurately predicted with a mutually exciting Hawkes process, however the topic models perform better for protest and violence.

The EM branching process formulation of the problem requires that the branching structure of the process be estimated simultaneously with the model parameters. Event cascades can then be detected using the triggering probability matrix p_{ij} from Section 2. In particular, we sample from the probabilities and then merge all events into clusters that are ancestors.

TABLE 5
NGI scores for Nigeria

	ME-HBTM	HBTM	ME-H
Protest	0.3159	0.2869	0.0853
Violence	0.4191	0.3700	0.2830
Territory	0.1549	0.0771	0.2662

In 5, we compare the UCI coherence [23] of the ME-HBTM model to the coherence of LDA (where the number of topics k for LDA is varied). Here we see that the ME-HBTM model has improved coherence, a measure of the tendency of words to co-occur at a greater frequency if they are in the same topic. This can be explained by the ME-HBTM construction, where parent-child events tend to share the same words in the model. Another advantage of the ME-HBTM is that the number of topics is automatically selected.

In Figure 6 we plot several of the largest clusters for each of the elections. In Kenya we detect several large clusters of events around the election debates, the election itself, protests, and a cluster corresponding to a supreme court petition regarding the validity of the election results. Three major topics arising in the Nigeria clusters center around the election, protests, and Boko Haram attacks. The self-excitation productivity parameter θ_{mm} of Twitter events for the Nigeria data set is close to 1 ($\approx .99$) yielding large, long-lasting clusters. For comparison, we display the results of LDA topic modeling in Figure 7. Unlike LDA, the HBTM is capable of detecting new clusters and automatically determines the number of clusters. In particular, there are over 100 topic-time clusters found in each data set by HTBM, which may be viewed as micro-topics in time compared to topics returned by LDA.

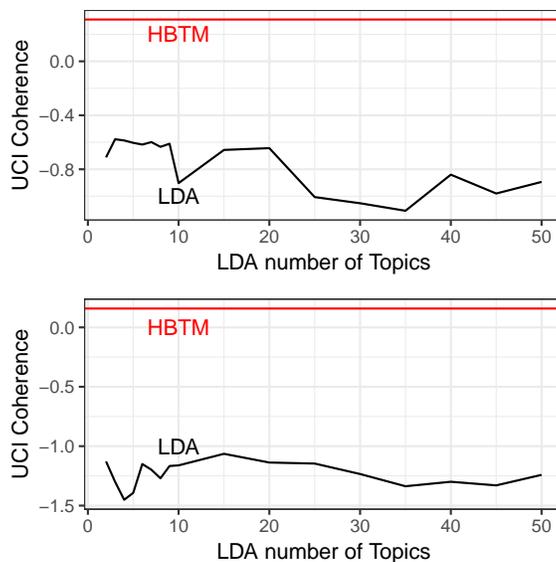


FIGURE 5. UCI coherence of HBTM vs. LDA for Nigeria (top) and Kenya (bottom) datasets.

4.3. *Goodness of fit.* In Equation 1 we make several assumptions on the triggering kernel, namely that the kernel is separable in topic space and time and that the distribution in time is exponential. In this sub-section we investigate the extent to which these assumptions may be valid. To do this we sample from the estimated branching probabilities p_{ij} in Equation 4. If the model is correctly specified, then sampled inter-event data $(t_i - t_j, m_i, m_j)$ will be realizations from the triggering kernel density. We can then assess separability and model fit based on this data.

First, we investigate separability in time and topic space. To reduce the dimension in mark space we apply principal component analysis to the document-term matrix and focus on the first principal component, which we denote by "PCA" in Figure 8. We then consider the inter-event pairs $(t_i - t_j, PCA_i - PCA_j)$ and conduct a Hoeffding test for dependence using the R package `testforDEP`. We find that in the case of Kenya we cannot reject the null hypothesis of independence, with a p-value of .08, though in the case of Nigeria we find some limited evidence of dependence among the two variables, with a p-value of .002. However, in Figure 8 we see that the correlation is low both in the case of Kenya (.004) and Nigeria (.022) and we do not find enough deviation from separability to warrant a more complex, non-separable model specification.

Next we investigate the goodness of fit of the exponential-in-time assumption for the triggering kernel. Again we sample from the estimated branching probabilities p_{ij} to yield realizations $t_i - t_j$ from the triggering kernel $\omega \exp(-\omega t)$ (under the assumption that the model is correctly specified). We note that in our multivariate Hawkes process model there are 16 cross-excitation kernels and that the data set size is small, $O(50)$, for some of the category-category pairs. We therefore restrict our attention to parametric models rather than consider non-parametric kernels that may be more appropriate for larger dataset sizes.

A common alternative choice in Hawkes processes to the exponential for the triggering kernel is a heavy tailed Pareto distribution and we compute the AIC for these two competing models for the diagonal components (self-excitation) in Table 6. Here we see that for protest, violence, and territory the one parameter exponential is favored over the two parameter Pareto. However, for Twitter the Pareto distribution provides a better fit. Nonetheless, in Figure 8 we compare the empirical histogram of the inter-event times for Twitter to a best fit exponential and note that the model provides a reasonable fit to the data.

	Exp AIC	Pareto AIC
protest	348.1	350.1
violence	209.5	210.8
Twitter	842.7	-510.3
territory	60.7	62.7

TABLE 6

Model fit of exponential vs. pareto diagonal (self-exciting) triggering kernels.

5. Significance and Impact. While previous point process studies of conflict and terrorism have modeled single event types, often in time only, we presented a framework that extends previous work to include cross-excitation within conflict and social media effects coupled with a topic model that allows for the incorporation of event descriptions. We believe that such a framework could find application not only in conflict studies, but in other scenarios where high dimensional information accompanies each event in a point process. For example, in the case of crime records databases often event descriptions are available but seldom used in space-time models of crime. The methodology outlined in this paper could be used to incorporate that information and to solve crime linkage problems where the goal is to match crimes committed by the same offender [25].

Our methodology also provides a tool for merging and summarizing coupled conflict-social media event datasets. We create and implement a framework to identify relevant conversations around elections thought to heighten tensions that expose underlying grievances between groups that have been and will be effected by those voted into power, and the perception of whether or not the electoral process was free and fair. We create an unbiased corpus by expanding our expert query with data-driven textual, social, and spatial dimensions that may have been omitted, and then using machine learners trained on human annotation to increase precision at scale. We couple data from these Tweets with conflict events, and perform event cascade detection where groups of events are identified that are both temporally, categorically (mark type), and topically related through the estimated branching process.

We illustrated this process using data from the time period around the 2015 Nigeria election and 2013 Kenya election, summarizing the regional instability and national conversations on the election during the time period. The ACLED event data captures all conflict or violent events, not just election-related violence. However, the model indicates that discourse on Twitter specifically related to the election likely drives protests more than violence and territory changes, reflected in the model’s assessed capabilities to capture a leading indicator for events. This is a significant finding for the

potential of national-level election conversations on social media to translate existing grievances into contentious action and or conflict.

In the future, it may be worthwhile to consider extensions of the HBTM model framework. In its current form the model ignores the potential dependency among various words, for example there is some probability that a word in a parent event is replaced by a synonym in the offspring event. Furthermore, there may be situations where a non-separable and/or non-parametric triggering kernel is appropriate. These will be directions for future research.

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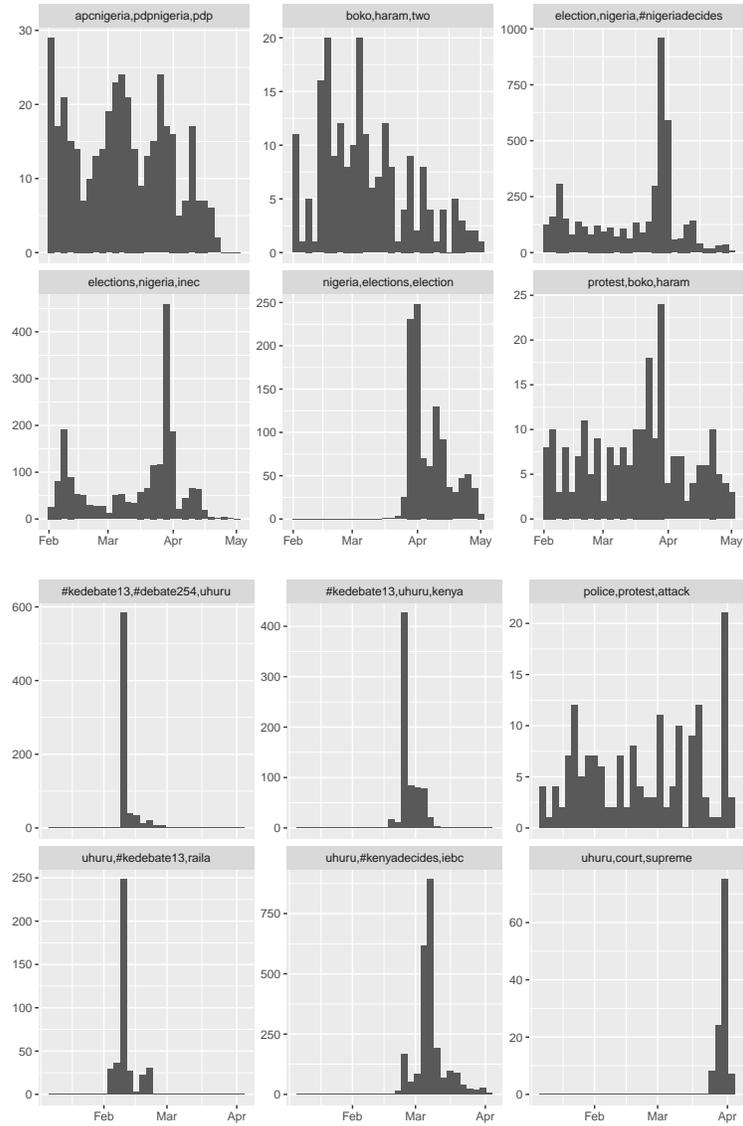


FIGURE 6. Histograms of event times of HBTM clustered events along with the three most frequent keywords words in each cluster for Nigeria (top two rows) and Kenya (lower two rows).

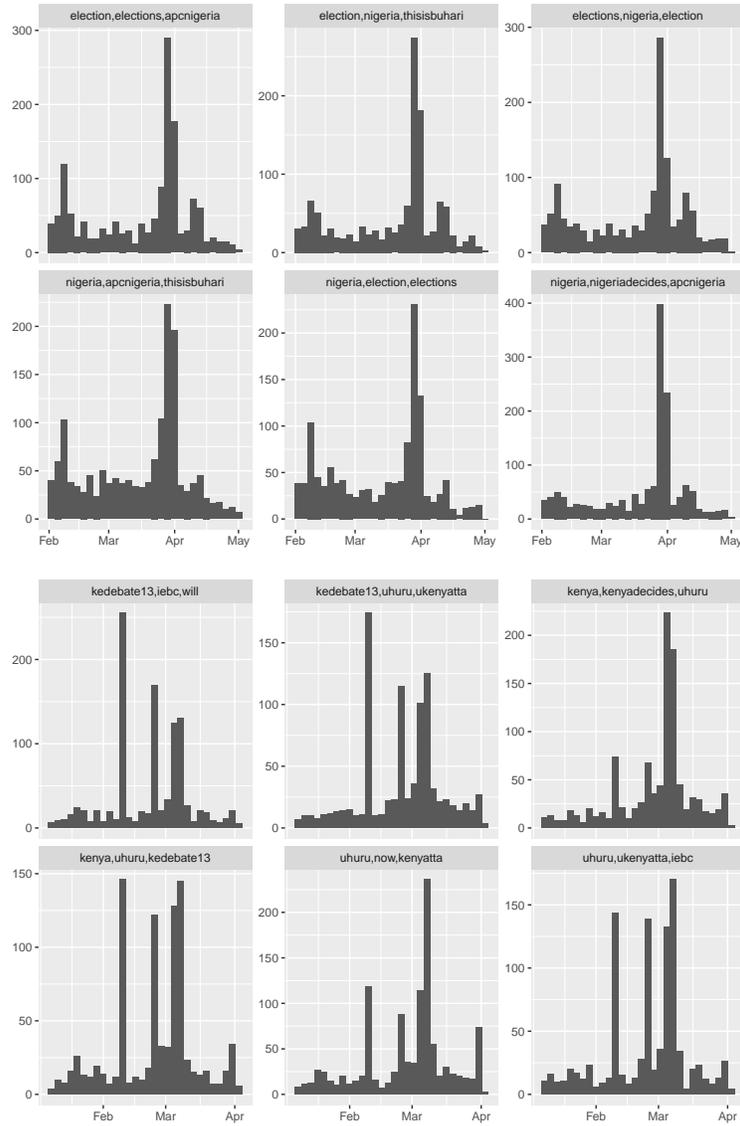


FIGURE 7. Histograms of event times of LDA clustered events along with the three most frequent keywords in each cluster for Nigeria (top two rows) and Kenya (lower two rows).

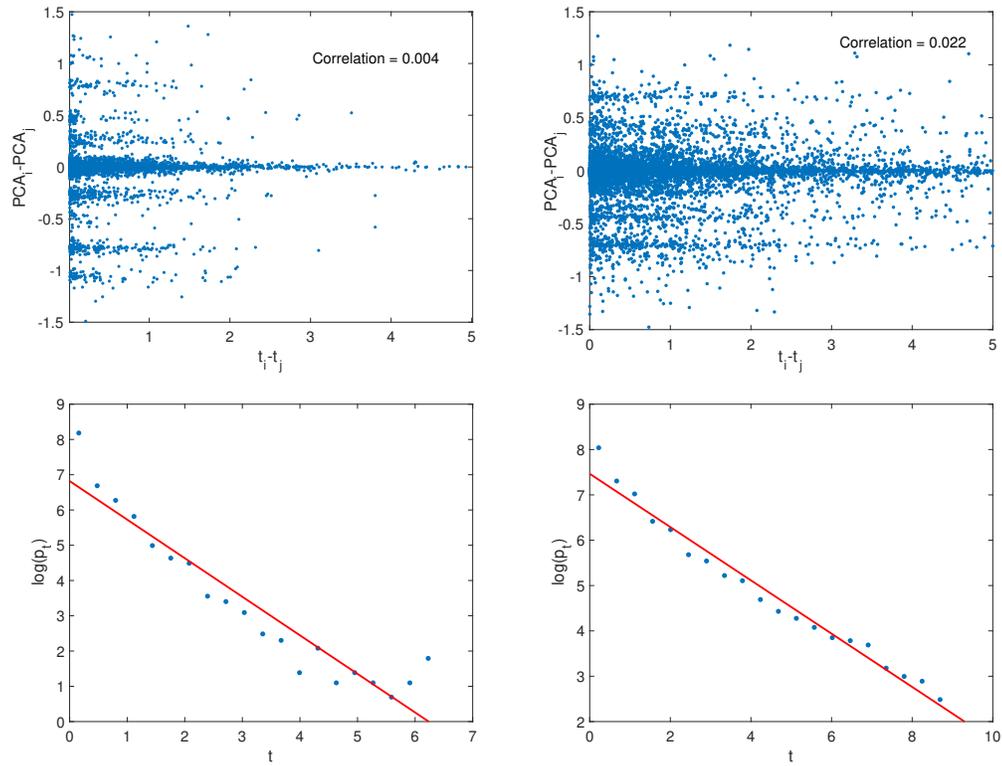


FIGURE 8. *Top: Scatter plot of the difference in mark first principal component versus difference in time of parent-offspring pairs for Kenya (left) and Nigeria (Right). Bottom: Best fit line to empirical histogram (log-scale) of sampled parent-offspring inter-event times $t_i - t_j$ for Kenya (left) and Nigeria (right).*